



Recent Advances in Regression Discontinuity Designs: Bridging Methods and Policy in Latin America

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Inequality

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Recent Advances in Regression Discontinuity Designs: Bridging Methods and Policy in Latin America

Sebastian Calonico UC Davis

2025 Ridge May Forum

- Today, I'll discuss recent advances in Regression Discontinuity (RD) designs, with a focus on connecting methodological innovations to policy evaluation practice in Latin America.
- I'll start by looking at how RD designs are currently used across the region:
 - What types of policies are being evaluated.
 - How researchers and practitioners implement RD in real-world settings.
 - What challenges commonly arise including issues with data, design choices, and interpretation.

- Understanding this context is important because Latin America has become an increasingly active setting for applied causal inference.
- Across areas such as education reforms, social protection programs, and health
 interventions, researchers and policymakers are applying a range of empirical strategies —
 including RD, RCTs, DiD, matching to evaluate the impacts of programs and policies.
- Among these approaches, RD designs are the natural choice when eligibility rules or thresholds create natural opportunities for identification.
- However, as RD applications have grown, so too have the complexities and limitations encountered in practice.

- In the second part of this talk, I'll turn to recent methodological developments that
 respond to these challenges innovations in RD design, estimation, and inference that
 aim to make applications more credible, robust, and informative for policymakers.
- My goal is to build a bridge between method and application, showing how these new tools can improve both the credibility and the usability of RD-based evaluations in the Latin American policy context.
- I'll focus on recent methodological developments in the estimation and inference of
 Heterogeneous Treatment Effects (HTE) in RD designs a topic of growing interest
 as policymakers and researchers aim to go beyond average effects and understand who
 benefits most, under what conditions, and by how much.
- These new tools aim to make RD analyses more informative and more policy-relevant.
- Let's begin by briefly reviewing the standard RDD setup.

Outline

Brief Intro to RDD

RD Policy Analysis in Latin America

Covariates in RDD

Heterogeneous RD Effects

Setup

Main Results

Mean Squared Error and Bandwidth Selection

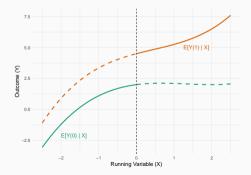
Inference

Empirical Application

Conclusion

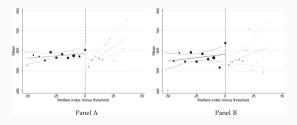
- Main goal: learn about the causal effect of a policy or intervention
- Units receive a **score**, X_i .
- Treatment is assigned based on the score and a *known* **cutoff** *c*:
 - given to units whose score is greater than the cutoff.
 - withheld from units whose score is less than the cutoff.
- Under continuity-type assumptions, the abrupt change in the probability of treatment assignment allows us to learn about the effect of the treatment on an **outcome** Y_i .

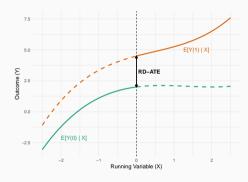
- Data: Random Sample (Y_i, T_i, X_i) , i = 1, 2, ..., n, with $Y_i = \begin{cases} Y_i(0) & \text{if } T_i = 0 \\ Y_i(1) & \text{if } T_i = 1 \end{cases}$
- Treatment: $T_i \in \{0,1\}$, $T_i = \mathbf{1}(X_i \ge c)$ X_i continuous

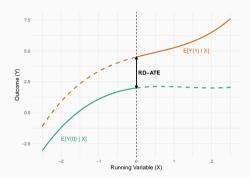


Example: Impact of **Social Health Insurance on Student Performance in Peru**: Carpio, Gomez & Lavado (2025), Health Economics.

- X_i: welfare index based on household characteristics.
- T_i : indicates whether households i is below a specific threshold
- c: welfare index cutoff.
- \bullet Y_i measures of student's performance.





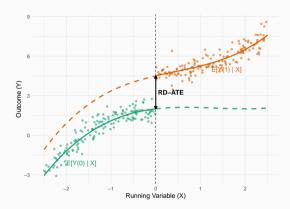


Average Treatment Effect at the cutoff:

$$\tau_{\text{SRD}} = \underbrace{\mathbb{E}[Y_{1i} - Y_{0i} | X_i = x_0]}_{\text{Unobservable}} = \underbrace{\lim_{\substack{x \downarrow x_0 \\ \text{Observable}}}}_{\text{Classified}} \mathbb{E}[Y_i | X_i = x] - \underbrace{\lim_{\substack{x \uparrow x_0 \\ \text{Observable}}}}_{\text{Observable}} \mathbb{E}[Y_i | X_i = x]$$

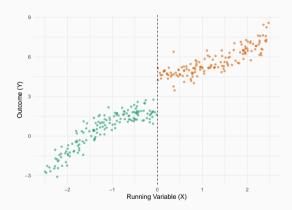
• We need to estimate $\lim_{x\downarrow x_0} \mathbb{E}[Y_i|X_i=x]$ and $\lim_{x\uparrow x_0} \mathbb{E}[Y_i|X_i=x]$.

Regression Discontinuity Design: Estimation



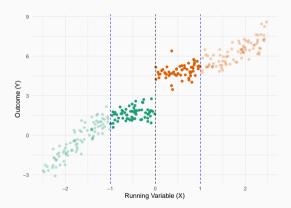
- Fundamental feature: no observations with score X_i exactly equal to the cutoff.
- Thus, extrapolation is unavoidable: we must rely on observations away from the cutoff.
- The unknown regression functions can be approximated by a polynomial function.

Regression Discontinuity Design: Estimation



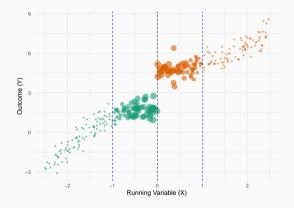
- Note: the approximation occurs at the cutoff, a **boundary point**.
- Global polynomial methods can lead to unreliable point estimators.
- Modern RD empirical work employs local (near the cutoff) polynomial methods.

Estimation: Local Polynomial Regression



- We only use observations between c h and c + h, where h > 0 is a **bandwidth**.
- Use data-driven (optimal) bandwidth selector. h_{MSE} : bias/variance trade-off.
- Statistical properties depend on the accuracy of the approximation.

Estimation: Local Polynomial Regression



• Observations closer to c receive more weight: **kernel** function K().

Estimation: Local Polynomial Regression



- Finally, we fit a polynomial of order p to construct point estimator $\hat{\tau}_n$ (optimal).
- Via joint or separate regressions at each side of the cutoff.

Choice of Bandwidth

- ullet Given p, find h to ensure optimal properties of the point estimator $\hat{ au_{RD}}$
- MSE-optimal plug-in rule:

$$MSE(\hat{ au_{RD}}) = Bias^2 + Variance \approx h^{2(p+1)}\mathcal{B}^2 + \frac{1}{nh}\mathcal{V}$$
 $h_{ ext{MSE}} = C_{ ext{MSE}}^{1/(2p+3)} \cdot n^{-1/(2p+3)}$ $C_{ ext{MSE}} = C(\mathcal{K}) \cdot \frac{\text{Var}(\hat{ au}_{ ext{SRD}})}{\text{Bias}(\hat{ au}_{ ext{SRD}})^2}$

ullet Key idea: trade-off bias and variance of point estimator $\hat{ au}$

$$\uparrow \; \mathsf{Bias}(\hat{\tau}) \Longrightarrow \downarrow \hat{h} \qquad \text{and} \qquad \uparrow \; \mathsf{Var}(\hat{\tau}) \Longrightarrow \uparrow \hat{h}$$

Inference

Given above steps, how do we make inference about τ ?

• Conventional approach: invert Wald test statistic under the null to form CI.

$$\mathcal{T}(h) = rac{\hat{ au}_
u(h) - au_
u}{\sqrt{\hat{\mathcal{V}}(\hat{ au}_
u(h))}} \sim^a \mathcal{N}(0, 1)$$

$$I_{ extsf{US}}(h) = \left[\hat{ au}_{
u}(h) - 1.96 imes \sqrt{\hat{\mathcal{V}}(\hat{ au}_{
u}(h))} \quad ; \quad \hat{ au}_{
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- It will only have correct asymptotic coverage if $nh^{2p+3} \rightarrow 0$, bandwidth "small enough".
- MSE-optimal bandwidth is "too large" $\mathbb{P}[\tau_{\nu} \in I_{US}(h_{MSE})] \not \to 1 \alpha$.
- Alternative: **undersmoothing** by selecting a bandwidth "smaller" than h_{MSE} for inference.

Modern Approaches to Inference

• Robust Bias Correction inference (Calonico, Cattaneo, Titiunik, 2014):

$$\mathcal{T}_{ exttt{RBC}}(h) = rac{\hat{ au}_{
u, exttt{BC}}(h) - au_{
u}}{\sqrt{\hat{\mathcal{V}}_{ exttt{BC}}(\hat{ au}_{
u, exttt{BC}}(h))}}, \qquad \hat{ au}_{
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u}(h) - h^{1+p-
u}\hat{\mathcal{B}}(h),$$

- Key feature: $\hat{\mathcal{V}}_{BC}(h)$ is an estimator of the variance of $\hat{\tau}_{\nu,BC}(h)$, not $\hat{\tau}_{\nu}(h)$.
- $I_{ exttt{RBC}}(h)$ remains valid under a wider set of bandwidth sequences $\mathbb{P}[au_
 u \in I_{ exttt{RBC}}(h_{ exttt{MSE}})] o 1 lpha$
- Calonico, Cattaneo & Farrell (2020): coverage error of $I_{RBC}(h)$ vanishes faster than $I_{US}(h)$.

Several Extensions

- Imperfect compliance: Fuzzy RD.
- Derivatives of the regression functions: Kink RD.
- Multi-cutoff RD.
- Multi-score RD / Geographic RD.
- Dynamic RD / Differences-in-Discontinuities.
- RD with Covariates.

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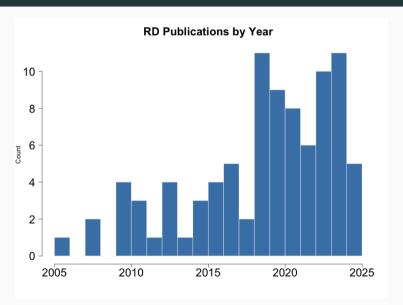
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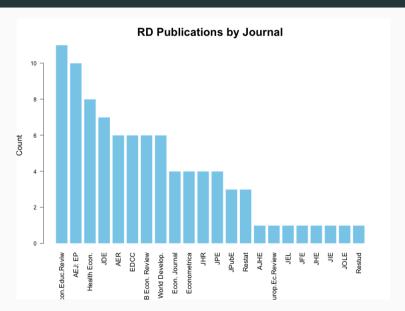
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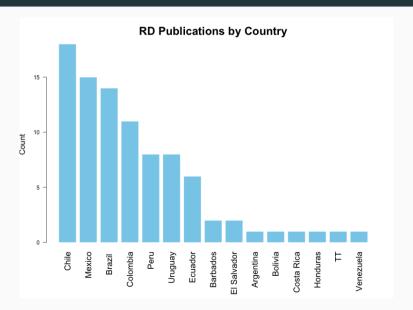
Empirical Application

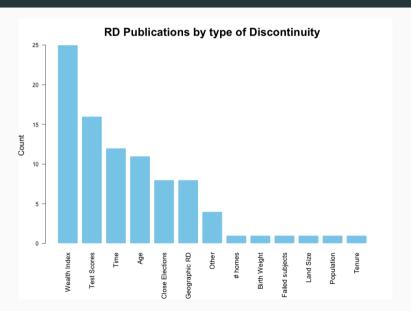
Conclusion

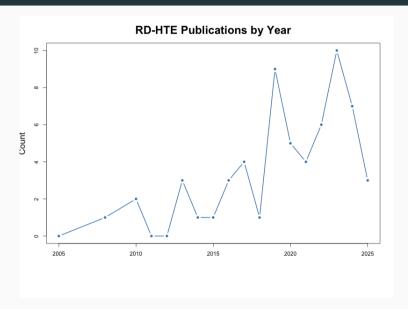
- Goal: understand how RD designs are currently used across the region:
 - What types of policies are being evaluated.
 - How researchers and practitioners implement RD in real-world settings.
 - What challenges commonly arise including issues with data, design choices, and interpretation.
- Review over the last 20 years of empirical papers using RDD as their main tool to evaluate policies in Latin America.
- Almost 30 journals, including top 5s and field journals.
- Found 90 articles.











Overview of empirical practice from recent RD publications.

- Most recent papers employ local polynomial methods, with data-driven (optimally) selected bandwidths.
- Need for extensions compared to the standard model: adjust for covariates, time trends, fixed effects, clustered standard errors, etc.
- This has implications for inference that are not typically accounted for.
- Growing interest in HTE analyses: mostly by subgroups, including discretize a continuous variable, and estimate separate or interacted RDDs.

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Covariates in RDD

- Standard RD model fits the outcome on the score alone.
- Sometimes we may want to augment it by including other covariates. Many possible goals:
 - 1. **Efficiency**: to capture finite sample variability as in RCTs.
 - 2. Alternative: to increase the plausibility of the design.
 - In general, covariate adjustment cannot be used to restore identification of standard RDD when treated and control observations differ systematically at the cutoff.
 - Continuity assumption of the potential outcomes is implausible.
 - Non-parametric framework no longer appropriate without further assumptions about dgp.
 - 3. A final use of covariates is to study **heterogeneous effects**.

- Heterogeneous treatment effects (HTE) is a key part of modern causal inference.
- Large methodological literature for estimation and inference on ATE, but comparatively little work has been done analyzing heterogeneity.
- This is particularly true for Regression Discontinuity Designs (RDD).
- RDD has become a widely used tool for causal inference in economics and social science, but most work focuses on ATE.
- This has lead to a proliferation of ad-hoc HTE analyses in applications.
- Goal: to provide practical and theoretically-grounded framework for HTE in RDD, yielding
 a unified methodology for empirical research.

Heterogeneous RD Effects via Covariates

The use of covariates in RDD to study heterogeneous effects can take several forms:

- 1. Multi-cutoff/Multivalued Treatments
- 2. Dynamic RD
- 3. Extrapolation based on CIA
- 4. Subgroup analysis

Heterogeneous RD Effects via Covariates

- 1. Multi-cutoff RD: Cattaneo, Keele, Titiunik, Vazquez-Bare (2016), Bertanha (2020).
 - Subgroups defined by the cutoff units are exposed to.
 - HTE: performing RD-ATE separately for each cutoff subgroup.
 - ATE at a particular cutoff for a subpopulation of units exposed to a different one.

2. Dynamic RD:

- Units are ranked and assigned to treatment repeatedly over time.
 - School districts referenda on bond issues, Cellini, Ferreira, & Rothstein (2010).
- Hsu & Shen (2024) recently provided a formal framework.
- 3. Extrapolation of TE based on CIA: Angrist, Rokkanen (2015), Keele et al (2015)
 - Effects far from the cutoff can be recovered with auxiliary pre-determined covariates.

Heterogeneous RD Effects via Covariates

Subgroup analysis

- Groups typically defined in terms of observed values of pre-determined covariates.
- The analysis changes the parameter of interest.

$$\tau_{Y|Z}(z) = E[Y_i(1) - Y_i(0)|X_i = c; Z_i = z]$$

- Typical approach: conduct RD analysis for each subgroup defined by $Z_i = z$, **separately**.
 - Example: TE patients for different age groups, e.g., 45-64 versus 65+.
- Very popular in practice, but no formal guidance on how to do it.

Heterogeneous RD Effects via Covariates

Subgroup analysis

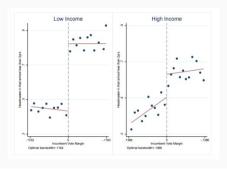
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- Typical approach: conduct RD analysis for each subgroup defined by $Z_i = z$, **separately**.
 - Example: TE patients for different age groups, e.g., 45-64 versus 65+.
- Very popular in practice, but no formal guidance on how to do it.
- Hsu & Shen (2019, 2021) propose formal tests for RD treatment effect heterogeneity among individuals with different observed pre-treatment characteristics.
- Reguly (2023), Alcantara et al (2024): tree- and forest- based methods for recovering RD heterogeneity, without specifying subgroups a priori.

Heterogeneous RD Effects in Applied Work

Akhtari, Moreira, Trucco (2022): political turnover in Brazil and public service provision.

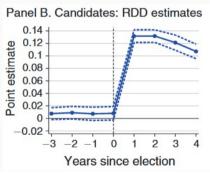


$$Y_{jmt+1} = (\alpha + \beta \mathbf{1}\{IVM_{mt} < 0\} + \delta \mathbf{1}\{IVM_{mt} < 0\} \times IVM_{mt} + \gamma IVM_{mt})$$

$$\otimes Charact_{jm} + X'_{jmt}\Lambda + \epsilon_{jmt}.$$

Heterogeneous RD Effects in Applied Work

Colonnelli, Prem, and Teso (2020): effect over time of supporting the winning party on probability of employment in the public sector.



(2)
$$y_{ikpmt} = \sum_{s=-3}^{+4} \beta_s Mayor_{pmt} \mathbf{1}(s = k) + \theta_k M V_{pmt} + \gamma_{kmt} + \epsilon_{ikpmt}.$$

Heterogeneous RD Effects in Applied Work

Overview of empirical practice from recent AEA publications.

| Paper | Outcome | Treatment | Running Variable | Heterogen | eity | | Estimation | | | | | Inference | |
|--------------------------|-------------------------------|--------------------------|------------------|-----------|------|-------|------------|-----------|-------|-------|-------------|-----------|---------------|
| | | | | Disc. | Time | Cont. | BW | HTE BW | Local | Joint | Covs/ FE | Robust | Cluster SE |
| Adams et al (2022) | Healthcare utilization | Financial Assistance | Poverty level | | Yes | | None | | | | | | |
| Akhtari et al (2022) | Municipal Bureaucracy | Political turnover | Close Elections | Yes | | | MSE | Yes | Yes | | Yes | | Yes |
| Asher, Novosad (2020) | Economic development | Road construction | Population size | Yes | | | MSE | | Yes | | Yes | | Yes |
| Brollo et al (2013) | Political corruption, quality | Government revenues | Population size | Yes | | Yes | None | | | Yes | Yes | | Yes |
| Dell (2015) | Drug-related violence | Drug enforcement | Close Elections | Yes | Yes | | Manual | | Yes | Yes | Yes | | Yes |
| Miralles, Leganza (2024) | Savings, Retirement behavior | Public pension benefits | Age | Yes | Yes | | MSE | | Yes | | Yes | Yes | |
| Han et al (2020) | Healthcare Utilization | Patient Cost-Sharing | Age | Yes | | | Manual | | Yes | | | | Yes |
| Huh, reif (2021) | Mortality, risky behaviors | Teenage Driving | Age | Yes | | | MSE | | Yes | | | Yes | |
| Jones et al (2022) | Irrigation Adoption, Profits | Access to water | Spatial Disc. | Yes | | | Manual | | Yes | | Yes | | Yes |
| Lindo et al (2010) | Academic Performance | Academic probation | Test Scores | Yes | | | Manual | | Yes | | Yes | | Yes |
| McEwan et al (2021) | Economics Major Choice | Higher letter grade | Test Scores | Yes | | | MSE | | Yes | Yes | Yes | Yes | |
| Miglino et al (2023) | Health outcomes | Financial Assistance | Age | Yes | Yes | | Manual | | Yes | | Yes | | Yes |
| Eleches, Urquiola (2013) | Academic Performance | Access to better schools | Test Scores | Yes | Yes | | Manual | | Yes | | Yes | | Yes |
| Shigeoka (2014) | Utilization, health | Patient Cost Sharing | Age | Yes | Yes | | Manual | | Yes | | Yes | | Yes |
| Silliman virtanen (2021) | Labor market returns | Vocational Education | Test Scores | Yes | Yes | | Manual | | Yes | | Yes | | Yes |
| Zimmerman (2019) | Better jobs, Income | Access to better schools | Test Scores | Yes | | | Manual | | Yes | Yes | Yes | | Yes |

- Most HTE analyses discretize a continuous variable, and estimate separate RDDs.
- Same bandwidth used for all analyses.
- Most inference is incorrect, unless willing to assume correctly specified model.

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Introduction

- Analyzing Heterogeneous Treatment Effects (HTE) is crucial in modern causal inference.
- While the ATE provides a broad measure of a treatment's effectiveness and informs general policy decisions, uncovering heterogeneity patterns is essential for evaluating substantive hypotheses and specific policy interventions.
- HTE can help understand fairness and differential impacts across subpopulations, develop targeted interventions, and optimize social policies.

Introduction

- RD design has become a prominent tool for causal inference, leading to extensive methodological advancements for identification, estimation, and inference for Average Treatment Effects (ATE) in various contexts.
- However, despite its widespread adoption, there is a notable lack of rigorous methods for studying heterogeneity, resulting in many ad-hoc empirical approaches.
- We address this gap by developing practical, theoretically-grounded tools for heterogeneity analysis in RD designs.

https://rdpackages.github.io/rdhte/

Treatment Effect Heterogeneity in Regression Discontinuity Designs*

Sebastian Calonico[†] Matias D. Cattaneo[‡] Max H. Farrell[‡] Filippo Palomba[‡]

Rocío Titiunik[‡]

March 26, 2025

Introduction '

- Extending RD estimation and inference to encompass HTE presents a high-dimensional, nonparametric challenge: theoretically feasible, but often impractical without restrictions.
- Consequently, empirical researchers typically use semiparametric models, and thus explore heterogeneity in limited dimensions.
- Following empirical practice, we study the most common approach for RD-HTE analysis: **local least squares regression with interactions** with pretreatment covariates.
- This model balances practical applicability with flexible (nonparametric) nature of RDD.

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- We consider a RD design with covariates **interacted** with treatment.
 - Popular approach in practice, but no formal results have been developed.
- Contribution 1: we show when this model correctly recovers useful causal objects.
 - When potential outcomes obey a linear functional coefficient model in covariates.
 - This structure holds without loss of generality for discrete variables.
 - Without such structure, linear interactions do not have a useful probability limit.
- Contribution 2: formalize how to conduct valid inference for HTE.
 - MSE expansion, central limit theorem, and standard error consistency.
 - We develop robust inference and optimal bandwidth selection methods.
 - Optimal bandwidth depends on specific inference target: subgroups, group differences, etc.

- Standard **Sharp RD** design based on a continuous running variable.
- Binary treatment $T_i \in \{0,1\}$, assigned according to $T_i = \mathbb{1}\{X_i \geq c\}$.
- In addition to (Y_i, T_i, X_i) , we observe a vector $\mathbf{W}_i \in \mathbb{R}^d$ of covariates.
- ullet Matching empirical practice, we assume $oldsymbol{W}_i$ are pre-treatment covariates.
- Usual parameter of interest in Sharp RD: ATE at the cutoff.

$$\tau = \mathbb{E}[Y_i(1) - Y_i(0) \mid X_i = 0]. \tag{1}$$

• The point estimate $\dot{\tau}$ is the coefficient on T_i in a (local, weighted) least squares regression:

$$\dot{Y}_i = \dot{\alpha} + \dot{\tau} T_i + \dot{\beta}_- X_i + \dot{\beta}_+ T_i X_i \tag{2}$$

- Calonico et al, 2019 added covariates to (2), but aiming for efficiency, not heterogeneity.
- Pre-treatment covariates Z_i enter linearly with "common" effect:

$$\widetilde{Y}_{i} = \widetilde{\alpha} + \frac{\widetilde{\tau}}{\tau} T_{i} + \widetilde{\beta}_{-} X_{i} + \widetilde{\beta}_{+} T_{i} X_{i} + \widetilde{\gamma}' \mathbf{Z}_{i}.$$
(3)

- Retain the same RD-ATE under "no RD effect on covariates" $\tilde{\tau} \to_{\mathbb{P}} \tau \left[\mu_{Z+} \mu_{Z-}\right]' \gamma_Y$
- Local linear projection, not assumption about functional form of regression functions.
- The exclusion of interactions with covariates is crucial for correct estimation of τ , but by construction does not allow for learning heterogeneity.

ullet One might try to recover the (local) CATE function fully flexibly in \mathbf{W}_i :

$$\tau(\mathbf{w}) = \mathbb{E}[Y_i(1) - Y_i(0) \mid X_i = 0, \mathbf{W}_i = \mathbf{w}]. \tag{4}$$

- $\tau(\mathbf{w})$: difference between two (d+1) dimensional nonparametric regressions evaluated at $(0,\mathbf{w})$.
- Challenging problem: curse of dimensionality, many tuning parameters.
- We aim for something more tractable.
- Thus, we focus on specifications with the same RD practical appeal: one dimensional nonparametric regression at a single point.
- Central question: to what extent, and under what assumptions, we can still obtain estimation and inference for heterogeneous causal effects.

• We study a local weighted least squares regression with full interactions.

$$\widehat{Y}_{i} = \widehat{\alpha} + \widehat{\tau} T_{i} + \widehat{\beta}_{-} X_{i} + \widehat{\beta}_{+} T_{i} X_{i} + \left(\widehat{\delta}'_{-} + \widehat{\delta}'_{+} T_{i} + \widehat{\rho}'_{-} X_{i} + \widehat{\rho}'_{+} T_{i} X_{i}\right) \mathbf{W}_{i},$$
 (5)

- Interpretation of coefficients changed.
- Without further assumptions on the dgp, the probability limit of the RD estimator has a generic best linear MSE predictor interpretation.
- We need to leverage extra structure to attach a causal interpretation.

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- We derive theoretical properties of the interacted regression fit.
- We first characterize exactly when the true CATE function is recovered.
- Next, we provide feasible inference with optimal tuning parameter selection.

 Besides standard RD regularity assumptions (modified to accommodate W), we impose two main conditions to obtain a causal interpretation of the estimands:

Assumption: Pre-determined covariates

At least in a neighborhood of the cutoff, \mathbf{W}_i is a *pre-treatment* or *exogenous* covariate,

$$\{\mathbf{W}_i(1),\mathbf{W}_i(0)\}\perp X_i$$

Assumption: CATE Structure

$$E[Y(t) \mid X = x, \mathbf{W} = \mathbf{w}] = \mu_t(x) + \lambda_t(x)'\mathbf{w}$$
 for $t \in \{0, 1\}$

where $\mu_t(x)$ and $\lambda_{t,k}(x)$, k = 1, ..., d are thrice continuously differentiable for all $x \in [x_l, x_u]$, $x_l < c < x_i$.

- Potential outcomes obey a linear-in-W structure (i.e. a local partially linear functional coefficient model).
- Assumption is placed on the potential outcomes, to match the two-sided nature of RDD.
- It does imply that the CATE function (defined at the cutoff) is:

$$\tau(\mathbf{w}) = \mathbb{E}[Y_i(1) - Y_i(0)|X_i = 0, \mathbf{W}_i = \mathbf{w}] = \tau(\mathbf{0}) + \delta'_{\star}\mathbf{w}$$

with
$$\tau(\mathbf{0}) = \mu_1(c) - \mu_0(c)$$
 and $\delta_{\star} = \lambda_1(c) - \lambda_0(c)$.

- The estimation of CATE function thus remains a one dimensional nonparametric problem.
- Bias can be removed up to the same order as learning the RD-ATE.
- Important for maintaining the empirical tractability of RD analyses.

- Good News: it holds without loss of generality for discrete covariates.
- Under these assumptions, $\widehat{\tau}(\mathbf{w}) \to_p \tau(\mathbf{w})$.
- ullet Thus, the coefficients $(\widehat{ au},\widehat{oldsymbol{\delta}}'_+)$, correctly capture the heterogeneous causal effects.
- Without it, the probability limit of the coefficients in (6) do not appear to have a causal interpretation.

$$\widehat{Y}_{i} = \widehat{\alpha} + \frac{\widehat{\tau}}{T_{i}} T_{i} + \widehat{\beta}_{-} X_{i} + \widehat{\beta}_{+} T_{i} X_{i} + \left(\widehat{\delta}'_{-} + \frac{\widehat{\delta}'_{+}}{\delta_{+}} T_{i} + \widehat{\rho}'_{-} X_{i} + \widehat{\rho}'_{+} T_{i} X_{i}\right) \mathbf{W}_{i}, \tag{6}$$

ullet For discrete covariates, the corresponding estimate of the CATE function $au(\mathbf{w})$ given by

$$\widehat{\tau}(\mathbf{w}) = \widehat{\tau} + \widehat{\boldsymbol{\delta}}'_{+}\mathbf{w}.$$
 (7)

- The coefficient on T_i no longer recovers the ATE, but rather $\hat{\tau} = \hat{\tau}(\mathbf{0})$, the CATE for the baseline group $\mathbf{w} = \mathbf{0}$.
- ullet The coefficients $\widehat{\delta}_+$ give the difference from the baseline group:
 - as discrete change for binary covariates, or
 - as slope of $\tau(\mathbf{w})$ with respect to continuous \mathbf{w} (under a functional coefficient assumption).

- Estimation and inference relies crucially on selecting a bandwidth *h* that appropriately localizes to the cutoff.
- We develop MSE-optimal bandwidth selection.
- This yields the best possible point estimates of heterogeneity.
- However, such a bandwidth yields invalid inference in general, and thus we also develop RBC inference methods.

- To characterize the optimal bandwidth we provide a MSE expansion of the coefficients $\widehat{\theta} = (\widehat{\tau}, \widehat{\delta}'_+)'$ around their population counterparts.
- Main difference: we now have several parameters of interest.
- Linear combinations of these yield the CATE function for different groups.
- Let **s** be the vector that selects the appropriate elements.
 - Binary Z_i : specific subgroup can be studied as $\widehat{\tau}(0) = \mathbf{s}'\widehat{\boldsymbol{\theta}}$ with $\mathbf{s} = (1,0)'$
 - Testing group differences: $\widehat{\tau}(1) \widehat{\tau}(0)$, which uses $\mathbf{s} = (-1,1)'$

MSE expansion:

$$\mathsf{MSE}\left[\mathsf{s}'\widehat{\theta} - \mathsf{s}' heta_\star
ight] = rac{1}{nh}\mathsf{V}_\mathsf{s} + h^4\mathsf{B}_\mathsf{s} + o_p\left((nh)^{-1} + h^4
ight).$$

- ullet V_s, B_s are the constant portions of the (asymptotic) variance and bias of $\mathbf{s}'\widehat{\boldsymbol{\theta}}$.
- Balancing the two terms yields the optimal bandwidth

$$h_{\mathbf{s}}^{\star} = \left(\frac{\mathsf{V}_{\mathbf{s}}}{\mathsf{B}_{\mathbf{s}}^{2}}\right) n^{-1/5} \tag{8}$$

- ullet h_s^\star depends on the specific inference target: each estimand has its own bandwidth.
- This is not necessarily practical. However, note:
 - The rate is not affected: this is a one-dimensional nonparametric problem (in X_i).
 - The constants are often less important empirically.
- Thus, we expect that results should not be too sensitive to the specific choice, allowing some flexibility in bandwidth selection.
- A common bandwidth could be the same used for obtaining the RD-ATE.

Inference

• Asymptotic distribution for inference on RD-HTE:

$$\sqrt{nh}\left(\widehat{\boldsymbol{\theta}}-\boldsymbol{\theta}_{\star}-h^{2}\mathsf{B}\right)
ightarrow_{d}\mathcal{N}(\mathsf{0},\mathsf{V}),$$

- h-MSE does not remove sufficient bias, rendering inference invalid.
- We propose RBC inference to deal with this problem. Two key ingredients:
 - 1. an estimate of h^2B is subtracted from the point estimate $\hat{\theta}$.
 - 2. standard errors adjusted to account for additional estimation error of the bias correction.
- We thus base inference on

$$\widehat{\boldsymbol{\theta}}_{\mathrm{rbc}} = \widehat{\boldsymbol{\theta}} - h^2 \widehat{\mathsf{B}}, \qquad \widehat{V} = \widehat{\mathrm{var}}(\widehat{\boldsymbol{\theta}}_{\mathrm{rbc}}).$$
 (9)

• Then, for any specific estimand defined by **s** and $h = h_{\mathbf{s}}^{\star}$:

$$\sqrt{nh} \widehat{V}^{-1/2} \left(\mathbf{s}' \widehat{oldsymbol{ heta}}_{
m rbc} - \mathbf{s}' oldsymbol{ heta}_{\star}
ight)
ightarrow_d \mathcal{N}(0,1).$$

- This result can be used to conduct inference on any estimand we discussed.
- For example, an asymptotic 95% confidence interval for $\tau(\mathbf{w})$ is given by

$$\widehat{ au}_{
m rbc} \pm 1.96 \sqrt{{\sf s}' \widehat{V} {\sf s}/(nh)}.$$

• Hypothesis tests can be constructed similarly. For example, $\tau(1) = \tau(0)$ test of no heterogeneity for a binary covariate.

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- To illustrate our methodological results, we revisit the analysis of Akhtari, Moreira, Trucco (2022) (AMT).
- Role of political turnover on public service provision by local governments in Brazil.
- Approach: compare outcomes in municipalities where the incumbent party barely loses (political turnover) to those where they barely win (no political turnover).
- Close election analysis, a canonical RD setting.
- Main findings: municipalities with a new party in office experience increases in bureaucracy, with new personnel appointed across multiple sectors and levels.
- Also, increases in the replacement rate of personnel in schools controlled by the municipal government, accompanied by lower test scores.
- We focus on heterogeneous effects on headmaster replacement (i.e., whether a headmaster is new to the school) by municipality income.

The analysis of AMT highlights two key methodological difficulties researchers have faced when studying heterogeneous treatment effects in RD.

- 1. They use an MSE-optimal bandwidth selection method to localize at the cutoff.
 - However, both estimation and inference rely solely on local linear fitting.
 - Invalid inference: standard errors from the linear fit are invalid when paired with the optimal bandwidth.
- 2. The heterogeneity variable (income) is continuous.
 - AMT compare municipalities above and below the median income level.

Table A.21: Political Turnover and Headmaster Replacement in Low- and High-Income Municipalities

| Outcome: | Headmaster is new to the school (as Headmaster) | | | | | | | | |
|-----------------------------|--|---------|---------|---------|---------|---------|--|--|--|
| Panel A | Low Income Municipalities (Below Median Income) | | | | | | | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | | | |
| $1\{IncumbVoteMargin < 0\}$ | 0.389 | 0.389 | 0.371 | 0.370 | 0.379 | 0.379 | | | |
| | (0.038) | (0.037) | (0.047) | (0.045) | (0.039) | (0.038) | | | |
| Observations | 6,703 | 6,703 | 4,294 | 4,294 | 6,447 | 6,447 | | | |
| R-squared | 0.151 | 0.154 | 0.160 | 0.167 | 0.156 | 0.159 | | | |
| Controls | No | Yes | No | Yes | No | Yes | | | |
| Clusters | 1073 | 1073 | 754 | 754 | 1030 | 1030 | | | |
| Mean Dep. Variable | 0.447 | 0.447 | 0.447 | 0.447 | 0.445 | 0.445 | | | |
| Using Bandwidth | 0.116 | 0.116 | 0.0700 | 0.0700 | 0.110 | 0.110 | | | |
| Optimal Bandwidth | 0.116 | 0.116 | 0.116 | 0.116 | 0.116 | 0.116 | | | |
| Panel B | High Income Municipalities (Above Median Income) | | | | | | | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | | | |
| $1\{IncumbVoteMargin < 0\}$ | 0.126 | 0.125 | 0.136 | 0.138 | 0.107 | 0.112 | | | |
| | (0.044) | (0.043) | (0.065) | (0.064) | (0.049) | (0.049) | | | |
| Observations | 5,809 | 5,809 | 3,114 | 3,114 | 4,560 | 4,560 | | | |
| R-squared | 0.050 | 0.051 | 0.030 | 0.032 | 0.045 | 0.046 | | | |
| Controls | No | Yes | No | Yes | No | Yes | | | |
| Clusters | 1220 | 1220 | 764 | 764 | 1048 | 1048 | | | |
| Mean Dep. Variable | 0.430 | 0.430 | 0.467 | 0.467 | 0.448 | 0.448 | | | |
| Using Bandwidth | 0.139 | 0.139 | 0.0700 | 0.0700 | 0.110 | 0.110 | | | |
| Optimal Bandwidth | 0.139 | 0.139 | 0.139 | 0.139 | 0.139 | 0.139 | | | |

- We first replicate their findings, adding RBC measures of uncertainty.
- Our point estimates are nearly identical (0.379 and 0.120 versus 0.389 and 0.126).
- Next, we incorporate valid inference, revealing that the treatment effect is statistically significant only for below-median income municipalities.
- We formally test that TE are the same across income groups, finding strong evidence against it.
- We also extend this discretized analysis by studying treatment effects by income quartiles and deciles, rather than a binary median split.
- Finally, we analyze income directly, as a continuous variable.

```
. rdhte Y X, covs_hte(i.W_qrt)

RD Heterogeneous Treatment Effects Estimation.

Cutoff c = 0
```

 Number of obs
 =
 206337

 Order est. (p)
 =
 1

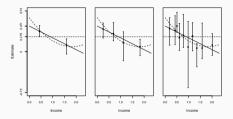
 Kernel
 =
 Triangular

 VCE method
 =
 hc3

Outcome: Y. Running variable: X.

| Group | Estimate | [95% Conf. | Interval] | P> z | Size | Bandwidth |
|---------|----------|------------|-----------|--------|-------|-----------|
| 1.W_qrt | . 41487 | 0.317 | 0.451 | 0.000 | 27414 | 0.142 |
| 2.W_qrt | .34048 | 0.318 | 0.450 | 0.000 | 18074 | 0.156 |
| 3.W_qrt | .08712 | -0.096 | 0.126 | 0.791 | 5372 | 0.098 |
| 4.W_qrt | .10157 | 0.029 | 0.188 | 0.008 | 9484 | 0.153 |

Application



- The continuous specifications highlight differences in income distribution within each group: top quartile has much wider range—an important finding for policy implications.
- These results further strengthen the original findings, showing that variation in treatment effects is more pronounced across the median than within the upper or lower groups.
- Ultimately, the decision of how to best capture heterogeneity is up to the researcher.
- The continuous approach, particularly when compared to a fine discretization, may yield substantial precision improvements, but requires an extra assumption.

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Conclusion

- Motivated from empirical practice, we study the statistical properties of the RD estimator when an interaction term between a covariate and the treatment is included.
- We showed how RD-HTE can be recovered using a tractable linear-interactions model.
- Our approach does not require high dimensional nonparametric regression and fits into current RD practice.
- By providing bandwidth selection and valid inference, we hope to bring the same rigor and standardization to heterogeneity in RD effects that currently exists for ATE.
- We have focused exclusively on Sharp RD (including Kink).
- Extending our work to other cases, particularly Fuzzy RD, would be a valuable next step.
- Robust inference, implementation details, replications:

https://rdpackages.github.io/